

ANALYSIS OF RICE CONSUMER PRICES IN INDONESIA, 2021–2025

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Abstract

Rice is a strategic staple food in Indonesia, and fluctuations in its consumer price have significant implications for food security and household welfare. This study aims to: (1) examine the trend and seasonal fluctuations of rice consumer prices in Indonesia during 2021–2025, (2) analyze the factors influencing rice price movements, and (3) project rice consumer prices for the subsequent two years. The study employs monthly secondary time-series data from January 2021 to December 2025. Descriptive analysis is used to identify price trends and seasonal patterns, while multiple linear regression is applied to assess the effects of inflation, rice production, rice import volume, and the rupiah-US dollar exchange rate on consumer rice prices. Price projections are conducted using linear trend analysis. The results indicate that rice consumer prices in Indonesia exhibit an overall upward trend accompanied by noticeable seasonal fluctuations. Empirical findings show that inflation has a negative and statistically significant effect on rice prices, while rice import volume and the rupiah-US dollar exchange rate have positive and significant effects. In contrast, rice production does not significantly influence consumer rice prices during the study period. Price projections suggest that rice consumer prices are expected to continue increasing over the next two years. The novelty of this study lies in its integrated analysis of trend behavior, seasonal fluctuations, and price projections using recent post-pandemic monthly data (2021–2025), providing updated empirical evidence on rice price dynamics in Indonesia. The findings offer practical implications for policymakers, particularly in designing effective price stabilization strategies, managing rice imports, and maintaining exchange rate stability to safeguard consumer purchasing power and national food security.

Keywords: Consumer Rice Prices, Exchange Rate, Inflation, Paddy Production, Rice Import Volume.

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INTRODUCTION

The substantial contribution of the agricultural sector to the national economy has led Indonesia to be recognized as an agrarian country. The agricultural sector plays a crucial role not only in food provision but also in employment absorption, income generation, and the maintenance of national economic stability. According to data from the Central Statistics Agency (BPS), the agricultural sector contributed approximately 11.31% to Indonesia's Gross Domestic Product (GDP) in 2024. One of the main pillars of the agricultural sector is food crops, particularly rice. Nearly 90% of the Indonesian population consumes rice on a daily basis, which is derived from paddy. The role of rice as a primary commodity is reflected in the high contribution of cereal-based foods to calorie and protein intake, which is significantly higher than that of other food groups. This condition positions rice not only as a staple food but also as a key determinant of national food security, whose production sustainability must be maintained. In addition to its high level of consumption, the relatively inelastic nature of rice demand reinforces the urgency of price control. Although per capita rice consumption has shown a gradual declining trend in recent years, the decrease has been relatively small and not proportional to the observed price changes. This indicates that rice remains an essential commodity that is difficult to substitute.

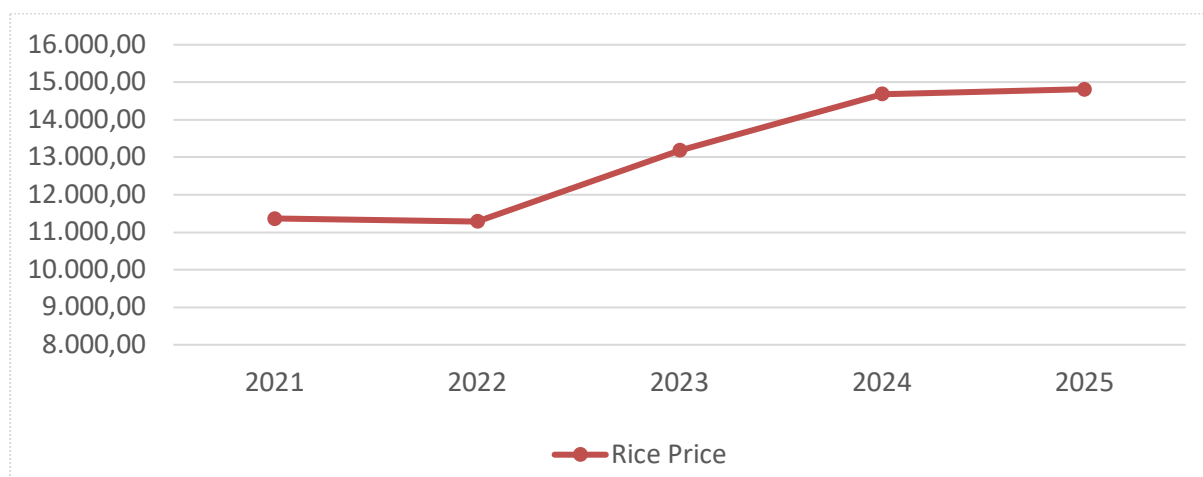


Figure 1. Development of Consumer Rice Prices in Indonesia, 2021-2025
Source: National Food Agency, 2021-2025

On the other hand, rice consumer prices exhibit more dynamic and fluctuating movements. As illustrated in Figure 1, rice consumer prices in Indonesia during the 2021–2025 period tend to show an upward trend despite experiencing fluctuations in certain periods. This persistent price increase reflects the presence of structural pressures within the food system, arising from supply-side constraints, distribution inefficiencies, and macroeconomic factors. Given the inelastic demand for rice, the upward price trend has the potential to increase household expenditure burdens and weaken consumers' purchasing power.

The Indonesian government does not rely solely on market mechanisms in regulating rice prices and availability, but also implements direct interventions to maintain price stability and ensure the availability of staple food commodities. One of the policy instruments is the establishment of a floor price for paddy to guarantee fair selling prices for farmers, along with a ceiling price for rice to protect consumers from sharp price increases (Rozci & Ristani, 2024). Furthermore, through Perum Bulog, the government conducts market operations by supplying rice to the domestic market at controlled prices in order to stabilize prices during periods of demand pressure or domestic supply shortages (Rahmasuciana & Darwanto, 2015).

This condition indicates that rice price formation is influenced not only by price policy instruments, but also by various other factors that operate simultaneously.

Previous studies have examined rice price dynamics using various approaches. Faladianti (2024) analyzed rice prices using an ARDL-ECM framework and found that exchange rates and imports influenced rice prices mainly in the short run, while domestic production showed limited effects. Damayanti et al. (2025) reported an increasing trend in rice prices and highlighted the role of macroeconomic variables and supply constraints. Anam et al. (2021) emphasized the role of inflation as an intervening variable between monetary factors and rice prices. However, most existing studies rely on annual or quarterly data, focus on specific regions, or emphasize long-run equilibrium relationships, thereby limiting their ability to capture short-term price fluctuations and seasonal patterns faced by consumers. This condition reveals a research gap in providing updated empirical evidence using recent monthly data that simultaneously examines price trends, seasonal fluctuations, and short-term determinants of rice consumer prices at the national level. Moreover, the post-pandemic period and climate disturbances have altered market dynamics, making earlier findings less applicable to current conditions.

This study hypothesizes that rice consumer prices in Indonesia exhibit an increasing trend and seasonal fluctuations during the 2021–2025 period. It is further hypothesized that inflation, paddy production, rice import volume, and the rupiah-US dollar exchange rate significantly influence changes in rice consumer prices in Indonesia over the same period. In addition, rice consumer prices are expected to continue showing an upward tendency in the two years following the observation period, namely 2026–2027.

The originality of this study lies in its integrated approach that combines trend analysis, multiple linear regression, and price projection using recent monthly data. This approach provides a more comprehensive and up-to-date understanding of rice price dynamics, offering valuable insights for policymakers in designing effective rice price stabilization and food security strategies.

METHODS

This study employs a quantitative approach with a time-series analysis design. The research is conducted at the national level, with the object of study being rice consumer prices in Indonesia during the period from January 2021 to December 2025. The use of monthly data aims to capture short-term price dynamics, long-term trends, and seasonal fluctuations. The data used in this study consist of monthly secondary time-series data obtained from official publications of the Central Statistics Agency (BPS), the National Food Agency, Bank Indonesia, the International Trade Centre (ITC), and other relevant institutions. The research variables include rice consumer prices (Y) as the dependent variable, and inflation (X_1), paddy production (X_2), rice import volume (X_3), and the rupiah-US dollar exchange rate (X_4) as independent variables. All data were screened and adjusted to ensure consistency in the observation period and time units prior to analysis.

Data analysis was conducted in several stages. The initial stage involved descriptive analysis to identify trends and seasonal fluctuations in rice consumer prices through price development graphs and monthly changes. Subsequently, the influence of the independent variables on rice consumer prices was analyzed using multiple linear regression, specified as follows:

$$Y = \alpha + \beta_1 X_1 + \beta_2 X_2 + \beta_3 X_3 + \beta_4 X_4 + \varepsilon$$

where:

Y = Rice consumer prices

α = Constant term (the value of rice prices when all independent variables are equal to zero)

$\beta_1 \dots \beta_4$ = Regression coefficients of the respective independent variables

$X_1 \dots X_4$ = Independent variables (inflation, paddy production, rice import volume, and the rupiah-US dollar exchange rate)

ε = Error term

To ensure that the regression model meets the criteria of unbiased and efficient estimation, several classical assumption tests were performed, including:

1. Normality Test

The normality test was conducted to examine whether the residuals are normally distributed. The Kolmogorov-Smirnov test was employed, where a significance value greater than 0.05 indicates normally distributed residuals.

2. Multicollinearity Test

The multicollinearity test aims to detect correlations among independent variables. This test was conducted by examining the Variance Inflation Factor (VIF) and tolerance values. The model is considered free from multicollinearity if $VIF < 10$ and tolerance > 0.10 .

3. Heteroskedasticity Test

The heteroskedasticity test was used to determine whether the residual variance is constant. The Glejser test was applied, with a significance value greater than 0.05 indicating the absence of heteroskedasticity.

4. Autocorrelation Test

The autocorrelation test was conducted to identify correlations among residuals across time periods. The Durbin-Watson test was employed, where the Durbin-Watson statistic falling between the lower and upper bounds ($dU < DW < 4 - dU$) indicates no autocorrelation.

After the regression model satisfied all classical assumption tests, hypothesis testing was performed to measure the magnitude of the influence of the independent variables on the dependent variable, including:

1. A simultaneous test (F-test) to assess the joint effect of all independent variables on the dependent variable.
2. Partial tests (t-tests) to examine the individual effect of each independent variable.
3. The coefficient of determination (R^2) was used to evaluate the explanatory power of the model in explaining variations in rice consumer prices.

Price projections for rice consumer prices for the two years following the observation period were conducted using trend analysis, specified as:

$$Y_t = a + bX_t$$

where:

Y_t = Trend value at a given period

a = Model constant

b = Trend slope coefficient

x = Time period index

t = Time

The projection results are used to provide an overview of future tendencies in rice prices. All data processing and analysis were conducted using SPSS and Microsoft Excel statistical software.

RESULTS AND DISCUSSION

Trends and Seasonal Fluctuations on Rice Consumer Prices

The presentation of rice consumer price data in a chronological order from January 2021 to December 2025 provides an initial overview of price movement patterns during the study period. The following section illustrates monthly rice price movements, which serve as a basis for analyzing the factors influencing the stability and changes in rice consumer prices.

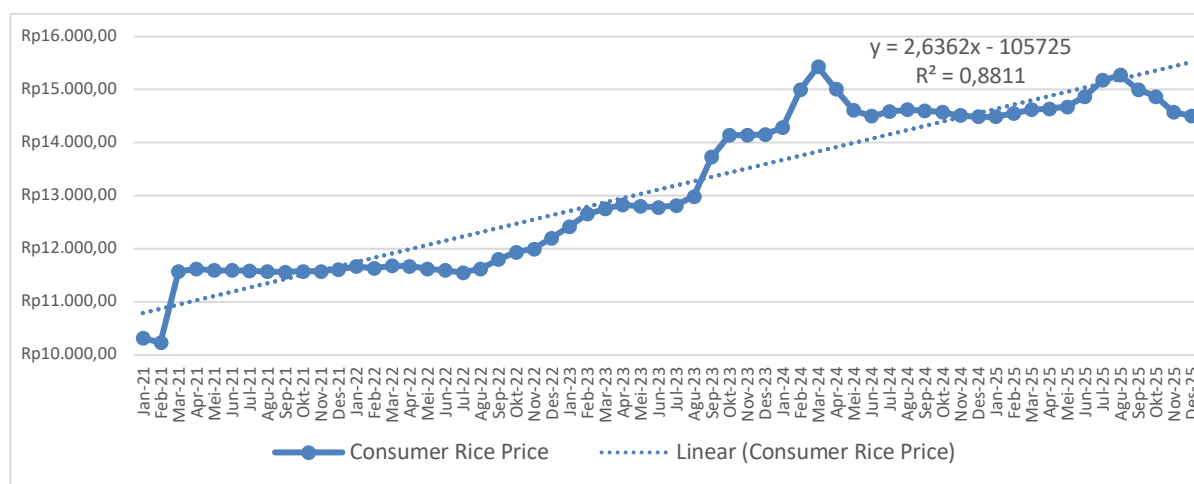


Figure 2. Trends and Seasonal Fluctuations in Rice Consumer Prices
Source: National Food Agency, 2025 (processed)

The results of the time-series analysis indicate that rice consumer prices in Indonesia during the 2021–2025 period experienced a consistent long-term upward trend, despite being accompanied by seasonal fluctuations. This pattern confirms that structural factors are more dominant than seasonal factors in shaping rice price dynamics. The sharp price increase during the 2023–2024 period, which did not fully return to pre-crisis levels, indicates the presence of price asymmetry, whereby prices tend to increase more rapidly than they decline after supply conditions improve.

From a food economics perspective, these findings reinforce the argument that rice exhibits inelastic demand characteristics, meaning that supply disruptions caused by extreme weather, distribution constraints, or external pressures are directly reflected in increases in consumer prices. This result is consistent with the findings of Rahman et al. (2020) and Damayanti et al. (2025), who state that production shocks and cost pressures tend to drive rice prices upward on a more permanent basis. The empirical contribution of this study lies in the use of recent monthly data, which reveals that the upward price trend is no longer temporary but has instead formed a new structural pattern in the Indonesian rice market.

Factors Influencing The Development of Rice Consumer Prices

This study aims to identify the factors influencing the development of rice consumer prices in Indonesia during the period from January 2021 to December 2025. The factors analyzed include inflation in Indonesia, paddy production, rice import volume, and the rupiah-US dollar exchange rate. The analysis in this study employs multiple linear regression. To ensure the validity of the regression results, classical assumption tests are conducted prior to the regression analysis.

Classic Assumption Tests

The Normality Test aims to find out whether the residual data in the regression model is distributed normally. The test was carried out using the One-Sample Kolmogorov-Smirnov Test, with Asymp results. Sig. (2-tailed). The results of the normality test are presented in the following Table 1.

Table 1. One-Sample Kolmogorov-Smirnov Test

		Unstandardized Residual
N		60
Normal Parameters ^{a,b}	Mean	.0000000
	Std. Deviation	514.87246573
Most Extreme Differences	Absolute	.099
	Positive	.099
	Negative	-.059
Test Statistic		.099
Asymp. Sig. (2-tailed) ^c		.200 ^d

Source: Processed Data, 2026

Based on the results of the normality test using the One-Sample Kolmogorov-Smirnov Test presented in Table 2, the asymptotic significance (Asymp. Sig.) value is 0.200, which is greater than 0.05. This indicates that the null hypothesis (H_0), stating that the data are distributed with a mean equal to zero, cannot be rejected at the 80% confidence level. Therefore, it can be concluded that the data in this study are normally distributed.

The multicollinearity test is conducted to determine whether there is a strong relationship among the independent variables in the regression model. Multicollinearity can cause instability in regression coefficients and thereby affect the accuracy of the analysis results. In this study, the multicollinearity test is performed by examining the Tolerance and Variance Inflation Factor (VIF) values presented in the Coefficients table. The results of the multicollinearity test are presented in Table 2 below.

Table 2. Coefficients^a in the Multicollinearity Test

Model		Collinearity Statistics	
		Tolerance	VIF
1	(Constant) Variable		
	Inflation	.996	1.004
	Rice Production	.935	1.070
	Rice Import Volume	.855	1.169
	Rupiah To Dollar Exchange Rate	.905	1.105

Source: Processed Data, 2026

As shown in Table 2, the tolerance values indicate that none of the independent variables has a tolerance value below 0.10, suggesting the absence of correlation among the independent variables. Similarly, the Variance Inflation Factor (VIF) values confirm this result, as none of the independent variables has a VIF value greater than 10. Therefore, it can be concluded that there is no multicollinearity among the independent variables in the regression model.

The heteroskedasticity test in this study is conducted graphically using a scatterplot between the Standardized Predicted Values and the Standardized Residuals. The distribution pattern of the points in the scatterplot is used as a basis to assess the presence or absence of heteroskedasticity. The graphical results of the heteroskedasticity test are presented in Figure 1 below.

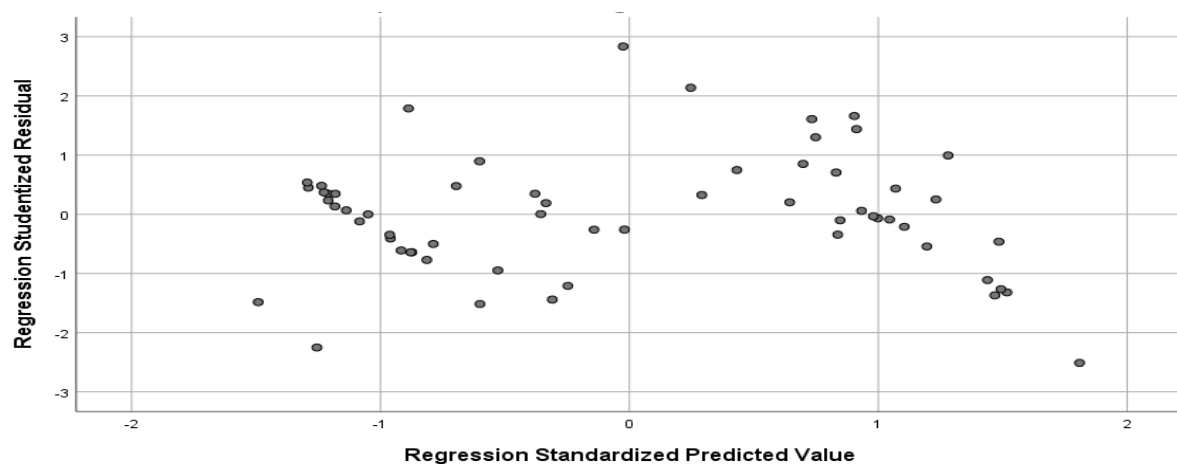


Figure 3. Scatterplot
Source: Processed Data, 2026

Figure 3 shows that the data points are randomly distributed and spread both above and below the zero value on the Y-axis. This indicates the absence of heteroskedasticity in the regression model. Therefore, the regression model is considered appropriate for predicting rice consumer prices based on the independent variables, namely inflation, paddy production, rice import volume, and the rupiah-US dollar exchange rate.

The autocorrelation test is conducted to determine whether there is a correlation between the error terms (residuals) across different observation periods in the regression model. In this study, the autocorrelation test is performed using the Durbin-Watson statistic reported in the Model Summary table. The results of the autocorrelation test are presented in Table 3 below.

Table 3. Model Summary^b in the Autocorrelation Test

Model	R	Std. Error of the Estimate	Durbin-Watson
1	.939a	533.26653	1.772

Source: Processed Data, 2026

The criterion for the Durbin-Watson autocorrelation test is determined by comparing the values $d_U < d_W < 4 - d_U$. These values are compared with the Durbin-Watson table at a 5% significance level. Based on these values, it can be concluded that the Durbin-Watson statistic ($d_W = 1.772$) is greater than the upper bound ($d_U = 1.727$) and less than $4 - d_U (2.273)$. Therefore, the model satisfies the criterion and indicates the absence of autocorrelation.

Hypothesis Testing

After the regression model has been confirmed to pass the classical assumption tests, hypothesis testing is conducted. The first test performed in the hypothesis testing is the coefficient of determination test. The coefficient of determination is used to explain the extent to which variations in the dependent variable can be explained by the independent variables, as measured by the Adjusted R-squared value.

Table 4. Model Summary^b in the Coefficient of Determination Test (R^2)

Model	R	R Square	Adjusted R Square	Std. Error of the Estimate
1	.939a	.881	.872	533.26653

Source: Processed Data, 2026

Based on the R-squared value of 0.872, it indicates that 87.2% of the variation in rice consumer prices can be explained by the independent variables, namely inflation, paddy production, rice import volume, and the rupiah-US dollar exchange rate, while the remaining proportion is influenced by other factors not included in the model.

The F-test represents the initial stage in assessing whether the estimated regression model is suitable for analysis. A model is considered appropriate if it is able to explain the joint effect of the independent variables (X) on the dependent variable (Y), as indicated by the coefficient of determination (Adjusted R²), which in this study amounts to 0.872. To examine the statistical significance of the Adjusted R² value, an F-test is subsequently conducted. The results of the F-test are presented in Table 5 below.

Table 5. Anova

Model		Sum of Squares	df	Mean Square	F	Sig.
1	Regression	115814450.448	4	28953612.612	101.816	.000 ^b
	Residual	15640525.702	55	284373.195		
	Total	131454976.150	59			

Source: Processed Data, 2026

It is known that the significance value for the simultaneous effect of inflation, paddy production, rice import volume, and the rupiah-US dollar exchange rate on rice consumer prices is $0.000 < 0.05$. In addition, the calculated F-statistic = 101.816 > F-table value = 2.54. Therefore, the null hypothesis (H₀) is rejected, indicating that the independent variables have a statistically significant simultaneous effect on the dependent variable.

The t-test is used to determine whether each independent variable partially has a significant effect on the dependent variable, thereby explaining the behavior of the independent variables in influencing the dependent variable. The results of the t-test are presented in Table 6 below.

Table 6. Coefficients^a in the t-Test

		Unstandardized Coefficients		Standardized Coefficients		
Model		B	Std. Error	Beta	t	Sig.
1	(Constant) Variable	-8816.635	1349.079		-6.535	.000
	Inflation	-199.222	50.482	-.184	-3.946	.000
	Rice Production	6.690	.000	.093	1.923	.060
	Rice Import Volume	2.599	.000	.293	5.830	.000
	Rupiah To Dollar Exchange Rate	1.426	.089	.787	16.093	.000

Source: Processed Data, 2026

The t-test results indicate the partial effects of each independent variable on rice consumer prices. Based on the statistical values presented in Table 6, the following findings are obtained:

1. Inflation (X₁) has a significance value of < 0.05 and an absolute t-statistic value of 3.946, which is greater than the t-table value of 2.004. Therefore, the null hypothesis (H₀) is rejected, indicating that inflation has a statistically significant partial effect on consumer rice prices. The estimated coefficient of -18.4 implies that a one percent increase in inflation leads to an 18.4% in rice prices, while a one percent decrease in inflation results in an 18.4% in rice prices. This finding is contrary to economic theory, which generally

suggests that inflation has a positive relationship with the prices of goods and services. In theory, an increase in inflation should be accompanied by an increase in rice prices. However, this result may be explained by the fact that rice is a strategic food commodity whose price is subject to government intervention and regulation. During periods of rising inflation, the government often intensifies food price stabilization policies, including market operations, the distribution of government rice reserves through food supply and price stabilization programs, and stricter monitoring of rice distribution channels. These measures are intended to maintain food affordability and prevent excessive increases in rice prices. As a result, rice prices may remain stable or even decline despite rising inflationary pressures in the broader economy. Therefore, the negative relationship between inflation and rice prices observed in this study suggests that rice prices are influenced not only by market mechanisms but also by government policies aimed at ensuring food security and price stability. This finding is consistent with the study conducted by Anam et al. (2021), which found that inflation had a significant negative effect on rice prices during the 2015–2019 period.

2. Paddy production (X_2) has a significance value of 0.060, which is greater than 0.05, and a t-statistic of 1.983, which is lower than the t-table value of 2.004. Thus, the null hypothesis (H_0) is accepted, indicating that paddy production does not have a statistically significant partial effect on rice consumer prices. This result is in line with the findings of Kurnia Fadilah et al. (2025), who reported that paddy production does not significantly affect rice prices. This condition occurs because rice prices tend to follow market mechanisms, so variations in paddy production do not necessarily translate into changes in rice price stability.
3. Rice import volume (X_3) has a significance value of 0.000, which is less than 0.05, and a t-statistic of 5.830, which is greater than the t-table value of 2.004. Therefore, the null hypothesis (H_0) is rejected, indicating that rice import volume has a statistically significant and positive partial effect on rice consumer prices. The estimated coefficient suggests that a one percent increase in rice import volume leads to an increase in rice prices by 29.3%, while a one percent decrease in import volume results in a price decline of 29.3%. This finding is consistent with Faladianti (2024), who reported that rice imports have a positive effect on rice prices in Indonesia in the short run. This condition arises because rice import policies in Indonesia are generally reactive, meaning imports are implemented when domestic prices have already increased or when domestic supply disruptions occur.
4. The rupiah-US dollar exchange rate (X_4) has a significance value of 0.000, which is less than 0.05, and a t-statistic of 16.093, which exceeds the t-table value of 2.004. Thus, the null hypothesis (H_0) is rejected, indicating that the exchange rate has a statistically significant and positive partial effect on rice consumer prices. Based on the estimated coefficient, a one percent increase in the rupiah-US dollar exchange rate leads to an increase in rice prices by 78.7%, while a decrease in the exchange rate results in a corresponding decline in rice prices by 78.7%. This result is consistent with the findings of Damayanti et al. (2025), who reported that the rupiah-US dollar exchange rate has a positive and significant effect on rice prices. This occurs because depreciation of the rupiah against the US dollar increases import costs, which ultimately contributes to higher rice prices in the domestic market.

Projection of Consumer Rice Prices

Trend and seasonal pattern analyses were conducted to identify the characteristics of consumer rice price movements during the 2021–2025 study period. Trend analysis was used to examine the general direction of rice price changes over time, while seasonal pattern analysis was employed to identify recurring monthly variations within a year. Seasonal

patterns were analyzed using the Seasonal Index method based on monthly consumer rice price data. The results of the Seasonal Index calculation are presented in Table 7 below.

Table 7. Seasonal Index of Consumer Rice Prices, 2021–2025

Month	Seasonal Index
Jan	96,06
Feb	97,46
Mar	100,45
Apr	100,02
May	99,32
Jun	99,37
Jul	99,94
Aug	100,49
Sep	101,43
Oct	102,05
Nov	101,58
Dec	101,84

Source: Processed Data, 2026

Based on the Seasonal Index calculation using consumer rice price data from 2021–2025, the highest seasonal index was recorded in October at 102.05, while the lowest was observed in January at 96.06. This indicates that rice prices in October tend to be approximately 2.05% above the annual average, whereas prices in January are about 3.94% below the annual average. Therefore, the difference between the highest and lowest seasonal indices is approximately 5.99%, suggesting the presence of a seasonal pattern in consumer rice prices.

However, the relatively small variation in seasonal index values indicates that the seasonal effect on consumer rice prices is weak. Rice prices tend to be lower at the beginning of the year, particularly in January–February, and relatively higher toward the end of the year, especially during September–December. Therefore, the movement of consumer rice prices during the study period appears to be influenced more by long-term trends than by seasonal factors. Therefore, a forecasting analysis was conducted to estimate the future movement of consumer rice prices. The forecast results for the period January 2026 to December 2027 are presented in Figure 4 below.

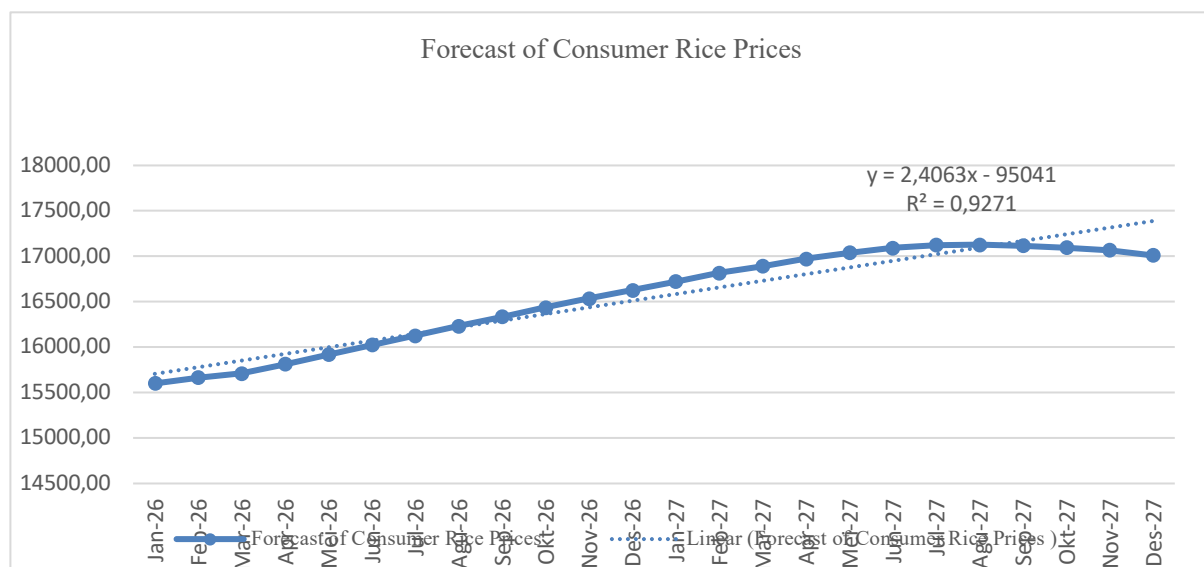


Figure 4. Forecast of Consumer Rice Prices for 2026–2027

Source: Processed Data, 2026

Based on Figure 4, the forecast results indicate that consumer rice prices are expected to increase gradually and remain relatively stable throughout the January 2026–December 2027 period. During the forecast horizon, rice prices are projected to reach a maximum increase of approximately 9.79% before experiencing a slight correction toward the end of the period. Nevertheless, the projected price at the end of the forecast period remains about 9.03% higher than the initial forecast value. The positive slope of the linear trend line further suggests that consumer rice prices are likely to continue increasing in the long run.

These projections are consistent with the trend and seasonal pattern analysis, which showed that consumer rice prices generally increased over time, while seasonal effects remained relatively weak, with a seasonal index difference of only 5.99%. This finding indicates that rice price movements are driven more by long-term trends than by seasonal factors. In addition, economic factors such as inflation, rice import volumes, and the exchange rate of the Indonesian rupiah against the US dollar may continue to exert upward pressure on consumer rice prices in the future.

CONCLUSION

Based on the results of the research and discussion, this study found that during the study period, from January 2021 to December 2025, the development of consumer rice prices showed a long-term upward trend accompanied by seasonal fluctuations. The regression results indicate that inflation, rice import volume, and the rupiah-US dollar exchange rate have a significant effect on consumer rice prices, whereas rice production does not have a significant impact. Furthermore, the forecasting analysis of consumer rice prices for the next two years, covering the 2026–2027 period, shows that consumer rice prices are expected to gradually increase with relatively stable movements.

The study recommends that the government strengthen rice price stability policies by controlling inflation, regulating imports, and maintaining rupiah exchange rate stability. Relevant institutions, such as the National Food Agency and Perum Bulog, should optimize their role in ensuring rice supply and price stability. Future research is encouraged to include additional variables that may influence consumer rice prices to provide a more comprehensive understanding of rice price dynamics in Indonesia.

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